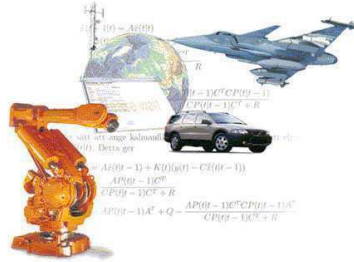


## Part 4 - Nonlinear system identification



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## The aim – Part 4

2(30)

The **aim in part 4** is to show how EM together with SMC and MCMC together with SMC can be used to solve challenging nonlinear system identification problems.

In other words, we will here make use of most of the building blocks introduced throughout the course in order to show how they can be combined to solve nonlinear system identification problems.



## Outline

3(30)

1. Computing ML estimates using EM and PS.
  - a) Derive the algorithm
  - b) Example - parametric Wiener model
2. Computing Bayesian estimates using particle MCMC (PMCMC)
  - a) Particle Marginal Metropolis Hastings (PMMH)
  - b) Example - semiparametric Wiener model



## ML problem formulation

4(30)

**Task:** Compute the ML estimate of the parameters  $\theta$  in the SSM

$$\begin{aligned}x_{t+1} | x_t &\sim f_{\theta}(x_{t+1} | x_t, u_t), \\ y_t | x_t &\sim h_{\theta}(y_t | x_t, u_t), \\ x_1 &\sim \mu_{\theta}(x_1),\end{aligned}$$

The ML estimate is obtained by solving the following optimisation problem,

$$\hat{\theta}^{\text{ML}} = \arg \max_{\theta} L_{\theta}(y_{1:N}),$$

where the log-likelihood function is given by

$$L_{\theta}(y_{1:N}) = \log p_{\theta}(y_{1:N}) = \sum_{t=1}^N \log p_{\theta}(y_t | y_{1:t-1})$$



The expectation maximisation (EM) algorithm computes ML estimates of unknown parameters in probabilistic models involving latent variables.

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**Algorithm 1** Expectation Maximization (EM)
 

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1. **Initialise:** Set  $i = 1$  and choose an initial  $\theta^1$ .
2. **While** not converged **do:**

(a) **Expectation (E) step:** Compute

$$Q(\theta, \theta^i) = E_{\theta^i} [\log p_{\theta}(Z, Y) | Y] = \int \log p_{\theta}(Z, Y) p_{\theta^i}(Z | Y) dZ$$

(b) **Maximization (M) step:** Compute  $\theta^{i+1} = \arg \max_{\theta \in \Theta} Q(\theta, \theta^i)$

(c)  $i \leftarrow i + 1$

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The **key property** rendering EM an appealing approach for computing maximum likelihood estimates in nonlinear SSMs is that the intermediate quantity  $Q(\theta, \theta^i)$  and its derivatives can be approximated arbitrarily well using particle smoothers.

EM provides a **strategy** for breaking down the problem into two manageable subproblems

1. A nonlinear state smoothing problem
2. A nonlinear optimisation problem

each of which can be handled using readily available algorithms.

The intermediate quantity  $Q(\theta, \theta^i)$  is approximated according to (using the particle smoother (FFBSi))

$$\widehat{Q}(\theta, \theta^i) = \widehat{I}_1(\theta, \theta^i) + \widehat{I}_2(\theta, \theta^i) + \widehat{I}_3(\theta, \theta^i),$$

where

$$\widehat{I}_1(\theta, \theta^i) = \frac{1}{N} \sum_{i=1}^N \log \mu_{\theta}(x_1^i),$$

$$\widehat{I}_2(\theta, \theta^i) = \frac{1}{N} \sum_{i=1}^N \sum_{t=1}^{T-1} \log f_{\theta}(x_{t+1}^i | x_t^i),$$

$$\widehat{I}_3(\theta, \theta^i) = \frac{1}{N} \sum_{i=1}^N \sum_{t=1}^T \log h_{\theta}(y_t | x_t^i).$$

Use a numerical nonlinear optimisation algorithm, e.g., BFGS. The gradient is computed according to

$$\nabla_{\theta} Q(\theta, \theta^i) = \nabla_{\theta} I_1(\theta, \theta^i) + \nabla_{\theta} I_2(\theta, \theta^i) + \nabla_{\theta} I_3(\theta, \theta^i),$$

and based on  $\widehat{Q}(\theta, \theta^i)$  it is straightforward to approximate these gradients according to,

$$\nabla_{\theta} I_1(\theta, \theta^i) \approx \frac{1}{N} \sum_{i=1}^N \nabla_{\theta} \log \mu_{\theta}(x_1^i),$$

$$\nabla_{\theta} I_2(\theta, \theta^i) \approx \frac{1}{N} \sum_{i=1}^N \sum_{t=1}^{T-1} \nabla_{\theta} \log f_{\theta}(x_{t+1}^i | x_t^i),$$

$$\nabla_{\theta} I_3(\theta, \theta^i) \approx \frac{1}{N} \sum_{i=1}^N \sum_{t=1}^T \nabla_{\theta} \log h_{\theta}(y_t | x_t^i).$$

**Algorithm 2** EM for nonlinear system identification

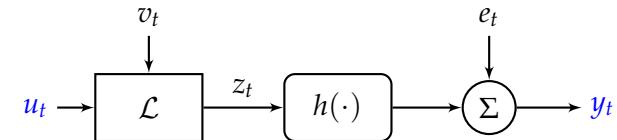
1. **Initialise:** Set  $i = 1$  and choose an initial  $\theta^1$ .
2. **While** not converged **do:**
  - (a) **Expectation (E) step:** Run a PF and a FFBSi PS and compute

$$\hat{Q}(\theta, \theta^i) = \hat{I}_1(\theta, \theta^i) + \hat{I}_2(\theta, \theta^i) + \hat{I}_3(\theta, \theta^i)$$

- (b) **Maximization (M) step:** Compute  $\theta^{i+1} = \arg \max_{\theta \in \Theta} Q(\theta, \theta^i)$  using an off-the-shelf numerical optimisation algorithm.
- (c)  $i \leftarrow i + 1$

Thomas B. Schön, Adrian Wills and Brett Ninness. **System Identification of Nonlinear State-Space Models.** *Automatica*, 47(1):39-49, January 2011. [pdf]

As an example we will study how to learn a **Wiener model**.



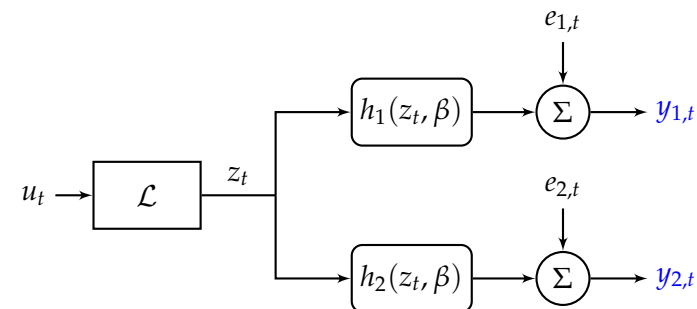
A Wiener model is a linear dynamical model ( $\mathcal{L}$ ) followed by a static nonlinearity ( $h(\cdot)$ ).

**Learning problem:** Find  $\mathcal{L}$  and  $h(\cdot)$  based on  $\{u_{1:T}, y_{1:T}\}$ .

Most of the existing work deals with special cases of the general problem. Typical restrictions imposed are:

- The nonlinearity  $h(\cdot)$  is assumed to be invertible.
- The measurement noise  $e_t$  is absent.
- The LGSS model is deterministic ( $v_t$  is absent).
- The LGSS model is stochastic, but  $v_t$  is assumed white.

Using EM + PS we do not have to impose any of these assumptions.



$$x_{t+1} = \begin{pmatrix} A & B \end{pmatrix} \begin{pmatrix} x_t \\ u_t \end{pmatrix}, \quad u_t \sim \mathcal{N}(0, Q),$$

$$z_t = Cx_t, \quad y_t = h(z_t, \beta) + e_t, \quad e_t \sim \mathcal{N}(0, R).$$

**Learning problem:** Find  $\mathcal{L}$  and  $\beta, r_1, r_2$  based on  $\{y_{1,1:T}, y_{2,1:T}\}$ .

The linear system ( $\mathcal{L}$ ) is given by

$$x_{t+1} = \begin{pmatrix} 1 & -0.9 \\ 1 & 0 \end{pmatrix} x_t + \begin{pmatrix} 1 \\ 0 \end{pmatrix} u_t,$$

$$z_t = \begin{pmatrix} 1 & 0.3 \end{pmatrix} x_t.$$

Complex poles implies a resonant system. The nonlinearities are a saturation and a dead zone, respectively,

$$h_1(z_t, \beta) = \begin{cases} \beta_1 & : z_t < \beta_1 \\ z_t & : \beta_1 \leq z_t \leq \beta_2 \\ \beta_2 & : z_t > \beta_2 \end{cases} \quad h_2(z_t, \beta) = \begin{cases} z_t - \beta_3 & : z_t < \beta_3 \\ 0 & : \beta_3 \leq z_t \leq \beta_4 \\ z_t - \beta_4 & : z_t > \beta_4 \end{cases}$$

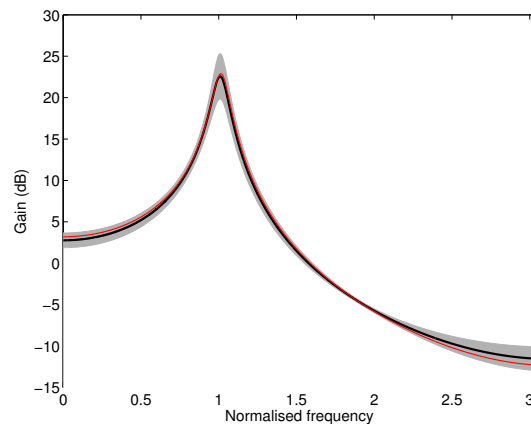
The measurements are given by

$$y_t = \begin{pmatrix} y_{1,t} \\ y_{2,t} \end{pmatrix} = \begin{pmatrix} h_1(z_t, \beta) \\ h_2(z_t, \beta) \end{pmatrix} + e_t, \quad e_t \sim \mathcal{N}\left(0, \begin{pmatrix} r_1 & 0 \\ 0 & r_2 \end{pmatrix}\right),$$

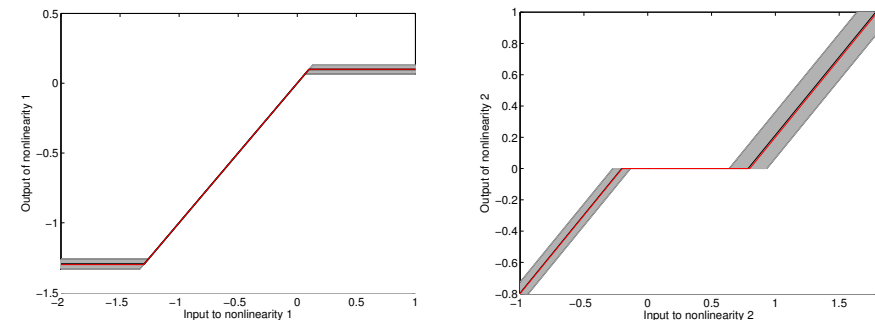
**The task** is to learn this model based on  $T = 1000$  measurements of the output (“blind” case),  $y_{1:1000}$ .

The input is chosen as  $u_t \sim \mathcal{N}(0, 1)$ . Initial values for the measurement variance are  $\hat{r}_1 = \hat{r}_2 = 0.1$ . The initial values for  $\hat{\eta}$  were chosen as  $\hat{\eta}_i = \frac{\eta_i^*}{10}$ , to reflect that they are unknown. The LGSS model is initialised via a subspace algorithm based on the measurements  $\{y_{1,1}, \dots, y_{1,T}\}$  from the dead zone nonlinearity.

Employ the EM alg. with  $N = 100$  particles. The algorithm was terminated after just 100 iterations. Plots below are based on 100 realisations of data.



Bode plot of estimated mean (black), true system (red) and the result for all 100 realisations (gray).



Estimated mean (black), true static nonlinearity (red) and the result for all 100 realisations (gray).

Adrian Wills, Thomas B. Schön, Lennart Ljung and Brett Ninness. **Identification of Hammerstein-Wiener Models.** *Automatica*, 2012. (In press) [\[pdf\]](#)

# Bayesian inference using Particle Markov chain Monte Carlo (PMCMC)



The **task** is to compute the pdf  $p(\theta | y_{1:T})$  for a model on the form

$$\begin{aligned}x_{t+1} | x_t &\sim f_t(x_{t+1} | x_t, u_t, \theta), \\ y_t | x_t &\sim h_t(y_t | x_t, u_t, \theta), \\ x_1 &\sim \mu(x_1, \theta), \\ \theta &\sim p(\theta).\end{aligned}$$

Can we set up an MCMC sampler to solve this problem?

Directly targeting  $\pi(\theta) = p(\theta | y_{1:T})$  is not possible, since  $p(\theta | y_{1:T})$  cannot be evaluated pointwise. Recall,

$$p(\theta | y_{1:T}) = \frac{p(y_{1:T} | \theta)p(\theta)}{p(y_{1:T})},$$

where it is not possible to pointwise evaluate the likelihood  $p(y_{1:T} | \theta)$  for the SSM above.



**Way forward:** Target  $\pi(\theta, x_{1:T}) = p(\theta, x_{1:T} | y_{1:T})$  instead. In order to understand why this works, note that

$$p(\theta, x_{1:T} | y_{1:T}) = \frac{p(x_{1:T}, y_{1:T} | \theta)p(\theta)}{p(y_{1:T})},$$

where

$$p(x_{1:T}, y_{1:T} | \theta) = \mu(x_1 | \theta) \prod_{t=1}^{T-1} f(x_{t+1} | x_t, \theta) \prod_{t=1}^T h(y_t | x_t, \theta).$$

This means that we can now evaluate the target density pointwise.



### We need one more thing for a working MCMC algorithm

How do we create a proposal distribution capable of proposing samples from relevant parts of the state space?

The state space  $\mathcal{X}^T \times \Theta$  is huge. Hence, it is key that it is explored in an efficient manner!

### How can this be done?

Make use of

- the model,
- and the observed measurements  $y_{1:T}$ .

Together they provide a lot of information about which parts of the state space that are most interesting.



Indeed, using SMC algorithms we can then turn the information in

- the model,
- and the observed measurements  $y_{1:T}$

into an efficient proposal distribution that captures what we know about where in the state space to propose new samples.

The result is a family of algorithms referred to as **particle MCMC (PMCMC)**.

The **fundamental idea underlying PMCMC** is to make use of an SMC sampler to construct a proposal for an MCMC sampler.

We will focus on the **particle Metropolis Hastings (PMH)** sampler in this course.

Notation:  $\mathbf{w}_t \triangleq \{w_t^1, \dots, w_t^N\}$ .

### Algorithm 3 Sequential Monte Carlo (SMC)

- 1. Initialise:** Sample  $x_1^i \sim Q_1(x_1)$  and set  $w_1^i = W_1(x_1^i)$ . Set  $t = 1$ .
- 2. For  $t = 2 : T$  do:**
  - (a) Resampling:**  $a_t^i \sim R(a_t | \mathbf{w}_{t-1})$ .
  - (b) Sample from the proposal kernel:**  $x_t^i \sim Q_t(x_t | x_{1:t-1}^{a_t^i})$  and set  $x_{1:t}^i = \{x_{1:t-1}^{a_t^i}, x_t^i\}$ .
  - (c) Weighting:**  $w_t^i = W_t(x_t^i, \tilde{x}_{1:t-1}^{a_t^i})$ .

Here:  $a_t^i$  to denote the **index** of the parent/ancestor at time  $t - 1$  of particle  $x_t^i$ .

Just like any algorithm that is used to generate random numbers there is an **underlying distribution** also for the SMC sampler that encodes the probabilistic properties of the involved stochastic variables.

The SMC sampler generates a realisation of the random variables  $\mathbf{X}_{1:T}$  and  $\mathbf{A}_{2:T}$ , where the pdf is

$$\psi(\mathbf{x}_{1:T}, \mathbf{a}_{2:T}) = \underbrace{\left\{ \prod_{i=1}^N Q_1(x_1^i) \right\}}_{\text{initialisation}} \left\{ \prod_{t=2}^T \prod_{i=1}^N \underbrace{R(a_t^i | \mathbf{w}_{t-1})}_{\text{Resampling}} \underbrace{Q_t(x_t^i | x_{1:t-1}^{a_t^i})}_{\text{Proposing new particles}} \right\}$$

and it is defined on the space  $\mathcal{X}^{TN} \times \{1, \dots, N\}^{(T-1)N}$ .

The SMC algorithm produce the following approximation of the target distribution

$$\hat{\pi}^N(x_{1:T}) = \sum_{i=1}^N W_T^i \delta_{x_{1:T}^i}(x_{1:T}).$$

We can use  $\hat{\pi}^N(x_{1:T})$  to produce samples of the state trajectory by sampling from

$$\begin{aligned} q^N(x_{1:T}) &= \mathbb{E}_\phi \left[ \hat{\pi}^N(x_{1:T}) \right] \\ &= \int \sum_{i=1}^N W_T^i \delta_{x_{1:T}^i}(x_{1:T}) \psi(\mathbf{X}_{1:T}, \mathbf{A}_{1:T-1}) d\mathbf{X}_{1:T} d\mathbf{A}_{1:T-1}. \end{aligned}$$

Note that this integration can never be carried out explicitly. Hence, we cannot evaluate this distribution for a specific sample, but it can be used to generate samples from it.

Use the following target distribution

$$\phi(\theta, \mathbf{x}_{1:T}, \mathbf{a}_{2:T}, k) \triangleq \frac{\pi(\theta, \mathbf{x}_{1:T}^k)}{N^T} \frac{\psi^\theta(\mathbf{x}_{1:T}, \mathbf{a}_{2:T})}{Q_1^\theta(x_1^{b_1^k}) \prod_{t=2}^T M_t^\theta(a_t^k, x_t^{b_t^k})}$$

Using this target we can now show that the acceptance probability is given by

$$a = \min \left( 1, \frac{\hat{p}(y_{1:T} | \theta^*) p(\theta^*) q(\theta[m-1] | \theta^*)}{\hat{p}(y_{1:T} | \theta[m-1]) p(\theta[m-1]) q(\theta^* | \theta[m-1])} \right)$$

1. **Initialise:** Set  $\theta[0]$  and run an SMC sampler targeting  $p(x_{1:T} | \theta[0], y_{1:T})$ , sample  $x_{1:T}[0] \sim \hat{p}^N(x_{1:T} | \theta[0], y_{1:T})$  and compute  $\hat{p}(y_{1:T} | \theta[0])$ .

2. **for**  $m = 1$  **to**  $M$  **do**

- (a) Sample  $\theta^* \sim q(\theta | \theta[m-1])$ .
- (b) Run an SMC sampler targeting  $p(x_{1:T} | \theta^*, y_{1:T})$ , sample  $x_{1:T}^* \sim \hat{p}^N(x_{1:T} | \theta^*, y_{1:T})$  and compute the  $\hat{p}(y_{1:T} | \theta^*)$ .
- (c) Compute the acceptance probability

$$a = \min \left( 1, \frac{\hat{p}(y_{1:T} | \theta^*) p(\theta^*) q(\theta[m-1] | \theta^*)}{\hat{p}(y_{1:T} | \theta[m-1]) p(\theta[m-1]) q(\theta^* | \theta[m-1])} \right)$$

- (d) With probability  $a$ , set the next state  $z[m] = \{\theta[m], x_{1:T}[m]\}$  of the Markov chain to  $\{\theta^*, x_{1:T}^*\}$  and  $\hat{p}(y_{1:T} | \theta[m]) = \hat{p}(y_{1:T} | \theta^*)$  and with probability  $1 - a$  set  $z[m] = \{\theta[m-1], x_{1:T}[m-1]\}$  and  $\hat{p}(y_{1:T} | \theta[m]) = \hat{p}(y_{1:T} | \theta[m-1])$ .

The PMMH sampler is a **standard MCMC sampler on a non-standard space**.

Put slightly differently, the PMMH sampler is a **standard MCMC sampler targeting a non-standard target distribution**.

It can be shown that the sequence of states and parameters  $\{\theta[m], x_{1:T}[m]\}_{m \geq 1}$  generated by the PMMH sampler will have  $p(\theta, x_{1:T} | y_{1:T})$  as its stationary distribution for any  $N \geq 1$  number of particles used in the PF!!

We can also derive a Particle Gibbs (PG) sampler with backward simulation (PG-BS). This is in fact what we used for the blind Wiener example mentioned in the introduction of Part 1.

General references for PMCMC:

- Christophe Andrieu, Arnaud Doucet and Roman Holenstein. **Particle Markov chain Monte Carlo methods**. *Journal of the Royal Statistical Society: Series B (Statistical Methodology)*, 72(3):269-342, June 2010. [\[pdf\]](#)
- Fredrik Lindsten and Thomas B. Schön. **On the use of backward simulation in the particle Gibbs sampler**. *Proceedings of the 37th International Conference on Acoustics, Speech, and Signal Processing (ICASSP)*, Kyoto, Japan, March 2012. [\[pdf\]](#)
- Fredrik Lindsten, Michael I. Jordan and Thomas B. Schön. **Ancestor Sampling for Particle Gibbs**. *Proceedings of Neural Information Processing Systems (NIPS)*, Lake Tahoe, NV, US, December, 2012. [\[pdf\]](#)

Using PMCMC for nonlinear system identification:

- Fredrik Lindsten, Thomas B. Schön and Michael I. Jordan. **A semiparametric Bayesian approach to Wiener system identification**. *Proceedings of the 16th IFAC Symposium on System Identification (SYSID)*, Brussels, Belgium, July, 2012. [\[pdf\]](#)

The **aim of this course** has been to provide an introduction to the theory and application of (new) computational methods for inference in dynamical systems.

The **key computational methods** we refer to are,

- Sequential Monte Carlo (SMC) methods (particle filters and particle smoothers) for nonlinear state inference problems.
- Expectation maximisation (EM) and Markov chain Monte Carlo (MCMC) methods for nonlinear system identification.

**Much interesting research remains** to be done in solving nonlinear inference/learning problems using SMC and/or MCMC methods!!

Feel free to contact me (now or later) in case you have questions or want to discuss things.

**Feedback (small and big)** on how to improve the lectures and the lecture notes are very welcome.

**Thank you for listening!!**